

## Chapter 7

50.  $X_1, \dots, X_n \perp \text{Unif}(0, \theta)$ 

a.  $f(x) = \frac{1}{\theta} \quad F(x) = \int_0^x \frac{1}{\theta} dt = \frac{x}{\theta}$

$$g_n(x) = n [F(x)]^{n-1} f(x) \\ = n \left[ \frac{x}{\theta} \right]^{n-1} \left( \frac{1}{\theta} \right) = \frac{n x^{n-1}}{\theta^n}, \quad x \in (0, \theta)$$

b.  $E[X_{(n)}] = \int_0^\theta \frac{n x^n}{\theta^n}$

$$= \frac{n}{\theta^n} \frac{1}{n+1} x^{n+1} \Big|_0^\theta = \frac{n}{n+1} \theta$$

c.  $E[X_{(n)}^2] = \frac{n \theta^2}{n+2}$

$$\Rightarrow V(X_{(n)}) = \frac{n \theta^2}{n+2} - \left( \frac{n}{n+1} \theta \right)^2$$

$$= \theta^2 \left( \frac{n}{n+2} - \frac{n^2}{(n+1)^2} \right)$$

d. (skip) soln. = 0 (easy)

e.  $R = X_{(n)} - X_{(1)}$

$$g_{1n} = \frac{n!}{(n-2)!} \left[ \frac{X_n}{\theta} - \frac{X_1}{\theta} \right]^{n-2} \frac{1}{\theta^2}$$

$$= \frac{n(n-1)}{\theta^n} (X_n - X_1)^{n-2}$$

X-Forms  $S = X_{(1)}$  temp variable

$$X_{(1)} = S \quad X_{(n)} = R + S$$

$$0 < S - r < S < \theta - r$$

$$0 < r < \theta - s$$

$$J = \begin{vmatrix} 0 & 1 \\ 1 & 1 \end{vmatrix} = -1 \quad |J| = 1$$

$$f_{RS} = g_{1n}(s, r+s)$$

$$= \frac{n(n-1)}{\theta^n} (r+s-s)^{n-2} = \frac{n(n-1)}{\theta^n} r^{n-2}$$

$$f_R = \int_0^{\theta-r} \frac{n(n-1)}{\theta^n} r^{n-2} ds = \frac{n(n-1)}{\theta^n} r^{n-2} (\theta-r)$$

f.  $E(R) = \int_0^\theta \frac{n(n-1)}{\theta^n} r^{n-1} (\theta-r) dr = \left( \frac{n-1}{n+1} \right) \theta$

## Chapter 8

2.  $E(X_i) = \int_{-1}^1 \frac{3}{2} x^3 dx = \frac{3}{8} x^4 \Big|_{-1}^1 = 0$

$$E(X_i^4) = \int_{-1}^1 \frac{3}{2} x^4 dx = \frac{3}{5} \underset{< \infty}{\cancel{b/c}} \quad E(x) = 0$$

By WLLN,

$$\bar{X}_n \xrightarrow{P} 0$$

4. By properties of the Poisson,

$$E(X_i) = \lambda \text{ and } V(X_i) = \lambda < \infty$$

So by WLLN,  $\bar{X}_n \xrightarrow{P} \lambda$ 

8.  $E(x) = \int_2^\infty \frac{2}{x} dx = 2 \ln x \Big|_2^\infty$

 $\Rightarrow \text{DNE} \quad V(x) \text{ cannot exist}$ 

WLLN does not apply

16.  $\mu = 12 \quad \sigma = 9 \quad n = 100$

By CLT,  $Y = \frac{\bar{X} - 12}{9/10} \sim N(0, 1)$

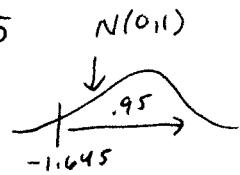
$$P(\bar{X} > 14) = P(Y > \frac{14-12}{9/10}) = P(Y > 2.22) = .0132$$

20.  $\mu, \sigma^2 = 4 \quad n = 50$

$$P(\text{Total} > 200) = .95 \quad N(0, 1)$$

$$P(\bar{X} > 4) = .95$$

$$-1.645 = \frac{4 - \mu}{2/\sqrt{50}}$$



$$\mu = 4.4653$$

8.24

$$Y_1 = \frac{\bar{X}_1 - \mu_1}{\sigma_1 / \sqrt{n_1}} \sim N(0, 1)$$

$$Y_2 = \frac{\bar{X}_2 - \mu_2}{\sigma_2 / \sqrt{n_2}} \sim N(0, 1)$$

$$\Rightarrow \bar{X}_1 - \mu_1 \sim N(0, \sigma^2 / n_1)$$

$$\bar{X}_2 - \mu_2 \sim N(0, \sigma^2 / n_2)$$

$$\Rightarrow (\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)$$

$$\sim N\left(0, \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right)$$

$$N(0, .3741657)$$

So,

$$P(|(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)| \leq .5)$$

$$P(|Z| \leq \frac{.5}{.3741657} = 1.336)$$

$$= .9082 - .0918 = .8164$$

8.34  $X \sim Exp(10)$   $n = 100$  $P(Y = 50 \text{ or more wait for } > 10 \text{ min})$ 

$$P(1 \text{ waits } > 10 \text{ min}) = P(X > 10) =$$

$$\int_{10}^{\infty} \frac{1}{10} e^{-x/10} dx = .3679 \text{ (rounded)}$$

$$Y \sim Bin(100, .3679) \quad P(Y \geq 50)$$

$$\text{check rule: } np = 36.79, n(1-p) \geq 10 \checkmark$$

$$\mu = 36.79 \quad \sigma = \sqrt{np(1-p)} = 4.822$$

$$P(Y \geq 50) \stackrel{CC}{=} P(W \geq 49.5) \\ = P(Z \geq \frac{49.5 - 36.79}{4.822}) =$$

$$P(Z \geq 2.64) = .0041$$

Use CC when  $\approx$  a discrete dist with a continuous dist. You don't need it working with  $\bar{X}$  b/c that's not discrete, but you do for counts/totals/etc.