

Moment Generating Functions

Let X be a simple random variable assuming the distinct values x_1, \dots, x_l with respective probabilities p_1, \dots, p_l . Its *moment generating function* is

$$(9.1) \quad M(t) = E[e^{tX}] = \sum_{i=1}^l p_i e^{tx_i}.$$

(See (5.19) for expected values of functions of random variables.) This function, defined for all real t , can be regarded as associated with X itself or as associated with its distribution—that is, with the measure on the line having mass p_i at x_i (see (5.12)).

If $c = \max_i |x_i|$, the partial sums of the series $e^{tX} = \sum_{k=0}^{\infty} t^k X^k / k!$ are bounded by $e^{|t|c}$, and so the corollary to Theorem 5.4 applies:

$$(9.2) \quad M(t) = \sum_{k=0}^{\infty} \frac{t^k}{k!} E[X^k].$$

Thus $M(t)$ has a Taylor expansion, and as follows from the general theory [A29], the coefficient of t^k must be $M^{(k)}(0)/k!$. Thus

$$(9.3) \quad E[X^k] = M^{(k)}(0).$$

Furthermore, term-by-term differentiation in (9.1) gives

$$M^{(k)}(t) = \sum_{i=1}^l p_i x_i^k e^{tx_i} = E[X^k e^{tX}];$$

taking $t = 0$ here gives (9.3) again. Thus the moments of X can be calculated by successive differentiation, whence $M(t)$ gets its name. Note that $M(0) = 1$.

Example 9.1. If X assumes the values 1 and 0 with probabilities p and $q = 1 - p$, as in Bernoulli trials, its moment generating function is $M(t) = pe^t + q$. The first two moments are $M'(0) = p$ and $M''(0) = p$, and the variance is $p - p^2 = pq$. ■

If X_1, \dots, X_n are independent, then for each t (see the argument following (5.10)), $e^{tX_1}, \dots, e^{tX_n}$ are also independent. Let M and M_1, \dots, M_n be the respective moment generating functions of $S = X_1 + \dots + X_n$ and of X_1, \dots, X_n ; of course, $e^{tS} = \prod_i e^{tX_i}$. Since by (5.25) expected values multiply for independent random variables, there results the fundamental relation

$$(9.4) \quad M(t) = M_1(t) \cdots M_n(t).$$

This is an effective way the sum S . The real interest so it is important to know from their moment generating

Consider along with (9.1) and suppose that $M(t) = N$ $M(t) \sim p_{i_0} e^{ix_{i_0}}$ and $N(t) \sim \zeta$ same argument now applies tively that with appropriate *the function (9.1) does uniquely*

Example 9.2. If $X_1, \dots, 0$ with probabilities p and successes in n Bernoulli trials generating function

$$E[e^{tS}] = ($$

The right-hand form shows distribution with mass $\binom{n}{k} p^k$ just established therefore yields

The *cumulant generating*

$$(9.5) \quad C(t)$$

(Note that $M(t)$ is strictly $(M')^2/M^2$, and since $M(0)$

$$(9.6) \quad C(0) = 0, \quad C$$

Let $m_k = E[X^k]$. The leading expansion of the logarithm is

$$(9.7) \quad C(t) =$$

Since $M(t) \rightarrow 1$ as $t \rightarrow 0$, this of 0. By the theory of series

This is an effective way of calculating the moment generating function of the sum S . The real interest, however, centers on the distribution of S , and so it is important to know that distributions can in principle be recovered from their moment generating functions.

Consider along with (9.1) another finite exponential sum $N(t) = \sum_j q_j e^{ty_j}$, and suppose that $M(t) = N(t)$ for all t . If $x_{i_0} = \max x_i$ and $y_{j_0} = \max y_j$, then $M(t) \sim p_{i_0} e^{ix_{i_0}}$ and $N(t) \sim q_{j_0} e^{ty_{j_0}}$ as $t \rightarrow \infty$, and so $x_{i_0} = y_{j_0}$ and $p_{i_0} = q_{j_0}$. The same argument now applies to $\sum_{i \neq i_0} p_i e^{ix_i} = \sum_{j \neq j_0} q_j e^{ty_j}$, and it follows inductively that with appropriate relabeling, $x_i = y_i$ and $p_i = q_i$ for each i . Thus the function (9.1) does uniquely determine the x_i and p_i .

Example 9.2. If X_1, \dots, X_n are independent, each assuming values 1 and 0 with probabilities p and q , then $S = X_1 + \dots + X_n$ is the number of successes in n Bernoulli trials. By (9.4) and Example 9.1, S has the moment generating function

$$E[e^{tS}] = (pe^t + q)^n = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k} e^{tk}.$$

The right-hand form shows this to be the moment generating function of a distribution with mass $\binom{n}{k} p^k q^{n-k}$ at the integer k , $0 \leq k \leq n$. The uniqueness just established therefore yields the standard fact that $P[S = k] = \binom{n}{k} p^k q^{n-k}$. ■

The *cumulant generating function* of X (or of its distribution) is

$$(9.5) \quad C(t) = \log M(t) = \log E[e^{tX}].$$

(Note that $M(t)$ is strictly positive.) Since $C' = M'/M$ and $C'' = (MM'' - (M')^2)/M^2$, and since $M(0) = 1$,

$$(9.6) \quad C(0) = 0, \quad C'(0) = E[X], \quad C''(0) = \text{Var}[X].$$

Let $m_k = E[X^k]$. The leading term in (9.2) is $m_0 = 1$, and so a formal expansion of the logarithm in (9.5) gives

$$(9.7) \quad C(t) = \sum_{v=1}^{\infty} \frac{(-1)^{v+1}}{v} \left(\sum_{k=1}^{\infty} \frac{m_k}{k!} t^k \right)^v.$$

Since $M(t) \rightarrow 1$ as $t \rightarrow 0$, this expression is valid for t in some neighborhood of 0. By the theory of series, the powers on the right can be expanded and